solutions 6

Solution 6.1. We know from Example 3.6 that

$$
p(x|y_{1:n}) = \mathcal{N}(x; \mu_p, \sigma_p^2),\tag{1}
$$

with

$$
\mu_p = \frac{\sigma_0^2 \sum_{i=1}^n y_i + \sigma^2 \mu_0}{\sigma_0^2 n + \sigma^2},
$$
\n(2)

$$
\sigma_p^2 = \frac{\sigma_0^2 \sigma^2}{\sigma_0^2 n + \sigma^2}.
$$
\n(3)

Below, we will zset $\mu_0 = 0$, $\sigma_0^2 = 1$ and $\sigma^2 = 1$ which simplifies example, i.e., we have the true mean

$$
\mu_p = \frac{\sum_{i=1}^n y_i}{M+1},\tag{4}
$$

which we will use for checking our code. In this exercise, I will give different parts of the solution separately to improve clarity.

1. In this part, we need to simulate our data.

```
1 \# simulate M data points
2 \mid M = 1003 | rng = np.random.default\_rng(25)4 \mid x = \text{rng.normal}(0, 1)5 | y = rng.normal(x, 1, M)
```
2. Secondly, we need to compute our true mean estimate as derived in [\(4\)](#page-0-0):

```
1 \mid # analytic posterior mean
2 mean_true = np.sum(y)/(M+1)3 print("Analytic posterior mean: ", mean_true )
```
3. (This is for parts 3-4 together) We will then implement SNIS with $\mu_q = 0$ and $\sigma_q^2 = 1$. For this, we directly define log densities.

```
1 \# define log prior
2 def logp(x):
3 return -x**2/2 - np.log(np.sqrt(2*np.pi))
4
5 # define log likelihood
6 def loglik(x, y):
7 return -(x-y)*2/2 - np.log(np.sqrt(2*p pip))8
9 \# define log proposal
10 def logq(x):
11 return -x**2/2 - np.log(np.sqrt(2*np.pi))12
13 def ESS(w):
14 return 1/np.sum(w**2)
15
16 \mid N = 1000017
18 x = rng.normal(0, 1, N) # sample from q(x)
```

```
19
20 \mid \text{log}W = \text{np}.\text{zeros}(N)21 for i in range(N):
22 logW[i] = np.sum(loglik(x[i], y)) + logp(x[i]) - logq(x[i])
23
24 \mid \text{log}_hat_W = \text{log}W - np.max(logW)
25
26 \leq w = np.\exp(\log\_hat_W) / np.\sum_{m \leq x} (log\_hat_W) ) # weights with
                                              log-trick
27 \mid w2 = np.\exp(\log W)/np.\sum(mp.\exp(\log W)) # weights without log-\frac{1}{2}trick
28
29 \# mean estimate
30 \vert mean = np. sum (w*x)31 | mean2 = np.sum (w2*x)32
33 print("Mean estimate (stable): ", mean)
34 print("ESS: ", ESS(w))
35 print("Mean estimate (unstable): ", mean2)
36 print("ESS: ", ESS(w2))
```
Solution 6.2. This is discussed in the lecture, so hopefully you will have more intuition about it. The first part of this exercise implements the SNIS.

```
1 import numpy as np
 2 import matplotlib .pyplot as plt
 3
4 def bar p(x): # implementing the density just for visualisation!
5 return np. exp(-x[0]**2/10 - x[1]**2/10 - 2 * (x[1] - x[0]**2)**2)6
7 \text{ def } q(x):
8 return np.exp(- x[0]**2/2 - x[1]**2/2) / (2 * np.pi)
9
10 \det logbar<sub>p</sub>(x):
11 return - x[0]*2/10 - x[1]*2/10 - 2 * (x[1] - x[0]*2)*212
13 def loglik(y, x, sig):
14 H = [1, 0]
15 return -(y - H @ x) * *2/(2 * sig **2) - np.log(sign * np.sqrt(2 * np.pi))
16
17 def logq(x):
18 return - x[0]**2/2 - x[1]**2/2 - np.log(2 * np.pi)19
20 def ESS(w):
21 return 1/np.sum(w**2)
22
23 |y = 124 \text{ sig} = 0.0525
26 \mid N = 1000027 \text{ rng} = \text{np.random.default\_rng}(25)28 \# sample from q
29 x = rng.normal(0, 1, (2, N)) # 2 x N matrix (2 dimensional, N samples)30
31 \# compute logW
32 \mid \text{logW} = \text{np}.\text{zeros}(N)33 for i in range(N):
```

```
34 | logW[i] = (loglik(y, x[:, i], sig)) + logbar_p(x[:, i]) - logq(x[:
                                               , i])
35
36 \# compute log hat W
37 \log hat W = logW - np.max(logW)
38 \vert w = np.\exp(\log_h at_W)/np.\sum_{m=1}^{N} (log_h at_W)39
40 # compute mean estimate
41 \vert mean = np.sum (w*x, axis=1)
42
43 # compute ESS
44 print("ESS: ", ESS(w))
```
Having obtained the weights w and the samples x, we can now resample. We will use the following code for resampling and plot the result.

```
1 \# resample N samples
2 | x_rresampled = np.zeros((2, N))
3 for i in range(N):
4 x<sup>-</sup> x<sup>-</sup> resampled[:, i] = x[:, rng.choice(N, p=w)]
5 # rng.choice chooses an index from 0 to N-1 with probability w
6
7 # plot resampled samples
8 | x_b = npu1inspace(-4, 4, 100)
9 | y_b = npulinspace(-2, 6, 100)
10 X_b, Y_bb = np.meshgrid(x_bb, y_bb)
11 |Z_b b = np{\text{.}zeros}((100, 100))12 for i in range(100):
13 for j in range(100):
14 Z_bbb[i, j] = bar_p([X_bb[i, j], Y_bb[i, j]])15 plt.contourf(X_bb , Y_bb , Z_bb , 100 , cmap='RdBu')
16 plt.scatter( x_resampled[0, :], x_resampled[1, :], s=10 , c='white')
17 plt.show ()
```
Note that, as explained in the lecture, the result makes sense. We had a 2D banana prior but only observed *x*₁ dimension with some noise as $y = Hx + \sigma W$ where $W \sim \mathcal{N}(0, 1)$ and since $H = [1, 0]$, this equals to $y = x_1 + \sigma W$ with small σ . It means that we can only know that the object in 2D will reside parallel to x_1 axis, according to our prior. Hence the samples from the posterior taking a vertical shape along this axis (would get even more vertical with smaller σ – watch the discussion in the lecture).